



Applications









Cramér-Rao bound for acoustic transfer function estimation

Giovanni Bologni (G.Bologni@tudelft.nl), Richard C. Hendriks, Richard Heusdens 31 May, 2024



On the menu today

Part 1 - Cramér-Rao bound for acoustic transfer function estimation

- 1. Parameter estimation & Cramér-Rao bound (CRB)
- 2. Case study ATF estimation

Part 2 - Acoustic transfer function estimation with inter-frequency correlation

- 1. Channel estimation algorithm
- 2. Experiments

Parameter estimation

A quantity θ needs to be estimated

Given a (random) model, what is best achievable accuracy on estimating unknown quantity?

Parameter estimation

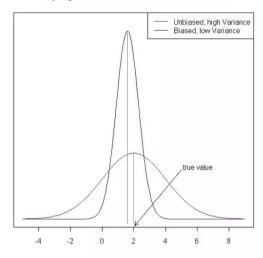
A quantity θ needs to be estimated

Given a (random) model, what is best achievable accuracy on estimating unknown quantity?

Best accuracy = minimum MSE = minimum variance (unbiased estimator)

$$\mathsf{MSE}(\theta, \hat{\theta}) = \mathsf{var}(\hat{\theta}) + \mathsf{bias}^2(\hat{\theta}, \theta).$$

Sampling Distributions of Estimated Parameters



Cramér-Rao bound

Under regularity assumptions on probability distribution $p(x; \theta)$,

$$\operatorname{var}(\hat{\theta}) \ge I^{-1}(\theta) = \frac{1}{-\mathbf{E} \left[\frac{\partial^2 \ln p(x;\theta)}{\partial \theta^2} \right]}.$$

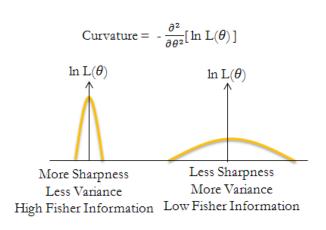
Pic: gaussianwaves.com

Cramér-Rao bound

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If PDF $p(x; \theta)$ is influenced by parameter more, estimation will be more accurate



Pic: gaussianwaves.com

Deterministic function of parameter

Suppose we want to estimate a function $g(\theta)$ of the parameter

Example sensor measures a quantity θ , but the instantaneous power θ^2 is needed: $q(\theta) = \theta^2$

In this case,

$$\operatorname{var}(\hat{\theta}) \geq \frac{\left(\frac{\partial g}{\partial \theta}\right)^2}{-\operatorname{\mathbf{E}}\left[\frac{\partial^2 \ln p(x;\theta)}{\partial \theta^2}\right]}.$$

Cramér-Rao bound – multiple parameters

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The Fisher information matrix (FIM) is the negative expected Hessian of the log-likelihood function:

$$\mathbf{I}_{\theta} = -\mathbf{E} \left[\nabla_{\theta} \nabla_{\theta}^{H} \ln p(\mathbf{x}) \right] = -\mathbf{E} \left[\nabla_{\theta}^{2} \ln p(\mathbf{x}) \right], \tag{1}$$

where the expectation is taken with respect to $p(\mathbf{x})$ and

$$[\nabla_{\theta} f]_i = \partial f / \partial \theta_i, \qquad [\nabla_{\theta}^2 f]_{ij} = \partial f / \partial \theta_i \partial \theta_j^*.$$

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The covariance matrix $\mathbf{R}_{\hat{\theta}}$ of any unbiased estimator $\hat{\theta}$ of θ satisfies $\mathbf{R}_{\hat{\theta}} \succeq \mathbf{I}_{\theta}^{-1}$.

 $A \succeq B$ means A - B is positive semidefinite with A and B Hermitian: $A = A^H$, $B = B^H$.

CRB – complex parameters

CRB described until now holds for real parameters. How to extend to complex parameters z?

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Two equivalent approaches:

- \triangleright Consider real Re(z) and imaginary part Im(z) separately (cumbersome)
- Consider complex number z and its conjugate z^* (also cumbersome \odot but generally easier)

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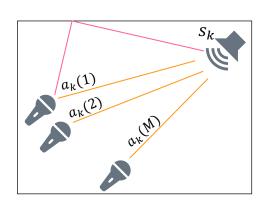
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Case study – channel estimation

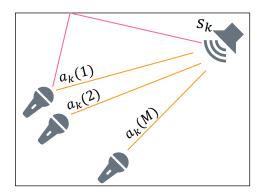


Signal model – frequency domain

Let a point source emit sound. The sound is measured by an array of M sensors. The received signal in the short-time Fourier transform (STFT) domain is

$$\mathbf{x}(l) = \mathbf{s}(l) + \mathbf{v}(l) = s(l) \mathbf{a} + \mathbf{v}(l) \in \mathbb{C}^M, \quad l = 1, \dots, L.$$

Our goal: recover transfer function a from noisy recording x



Deriving CRB for channel estimation

$$\operatorname{var}(\hat{\theta}) \ge I^{-1}(\theta) = \frac{1}{-\mathbf{E} \left\lceil \frac{\partial^2 \ln p(x;\theta)}{\partial \theta^2} \right\rceil}.$$

Likelihood function

Collect IID measurements in data matrix X

Assume noise v is complex circular Gaussian process

Unknown parameters $\theta = [\mathbf{a}^T \mathbf{a}^H]^T \in \mathbb{C}^{2M}$

Conditional likelihood is

$$p(\mathbf{X}; \theta, s(l)) = \frac{1}{|\pi \mathbf{R}|^L} \exp\left(-\sum_{l=1}^L (\mathbf{x}(l) - s(l)\mathbf{a})^H \mathbf{R}^{-1} (\mathbf{x}(l) - s(l)\mathbf{a})\right),$$

Log-likelihood and its derivatives

Define log-likelihood

$$L(\theta) = \ln p(\mathbf{X}; \theta) = -L \ln |\pi \mathbf{R}| - \sum_{l=1}^{L} (\mathbf{x}(l) - s(l)\mathbf{a})^{H} \mathbf{R}^{-1} (\mathbf{x}(l) - s(l)\mathbf{a})$$

We then have

$$\nabla_{\mathbf{a}^*} L(\theta) = \mathbf{R}^{-1} \sum_{l=1}^{L} (s(l)^* \mathbf{x}(l) - |s(l)|^2 \mathbf{a})$$

$$\nabla_{\mathbf{a}} L(\theta) = (\nabla_{\mathbf{a}^*} L(\theta))^*$$

$$- \mathbf{E} \left[\nabla_{\mathbf{a}^*} \nabla_{\mathbf{a}^*}^H L(\theta) \right] = - \mathbf{E} \left[\nabla_{\mathbf{a}^*} \nabla_{\mathbf{a}}^T L(\theta) \right] = \mathbf{E} \left[\mathbf{R}^{-1} \sum_{l=1}^{L} |s(l)|^2 \right] = E_s L \mathbf{R}^{-1}$$

$$- \mathbf{E} \left[\nabla_{\mathbf{a}^*} \nabla_{\mathbf{a}}^H L(\theta) \right] = \mathbf{0}$$

Fisher information matrix

With this, the Fisher information matrix is

$$\mathbf{I}_{\theta} = \begin{bmatrix} E_s L \mathbf{R}^{-1} & \mathbf{0} \\ \mathbf{0} & E_s L \mathbf{R}^{-*} \end{bmatrix}. \tag{1}$$

The block-diagonal matrix can be easily be inverted, leading to

$$\mathbf{I}_{\theta}^{-1} = \begin{bmatrix} \frac{1}{E_s L} \mathbf{R} & \mathbf{0} \\ \mathbf{0} & \frac{1}{E_s L} \mathbf{R}^* \end{bmatrix}$$
 (2)

Variance is finally bounded as:

$$\operatorname{var}(\hat{a}_i) \ge \frac{[R]_{ii}}{F_{i}L}, \quad i = 1, \dots, M. \tag{3}$$

Deterministic function of parameter

ATFs are often estimated in relation to a reference microphone r, as in $g(\theta) = g(\mathbf{a}, \mathbf{a}^*) = \mathbf{a}/a_r$. In this case,

$$\mathbf{R}_{g(\theta)} - (\nabla_{\theta} \mathbf{g}) \mathbf{I}_{\theta}^{-1} (\nabla_{\theta}^{H} \mathbf{g}) \ge 0, \tag{1}$$

where $\mathbf{R}_{g(\theta)}$ is the covariance matrix of $g(\theta)$. Choosing r=1, Jacobian is

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$$\nabla_{\theta} \mathbf{g} = \begin{bmatrix} \nabla_{\mathbf{a}} \mathbf{g} & \nabla_{\mathbf{a}^*} \mathbf{g} \end{bmatrix} \tag{2}$$

$$= \begin{bmatrix} 0 & 0 & 0 & 0 & 0 \\ -\theta_2/|\theta_1|^2 & 1/\theta_1 & 0 & \dots & 0 \\ -\theta_3/|\theta_1|^2 & 0 & 1/\theta_1 & \dots & 0 & \mathbf{0}_{M \times M} \\ \vdots & & \ddots & \vdots & \\ -\theta_M/|\theta_1|^2 & 0 & \dots & 0 & 1/\theta_1 \end{bmatrix},$$
(3)

where $[\nabla_{\theta} \mathbf{f}]_{ij} = \partial f_i / \partial \theta_j$.

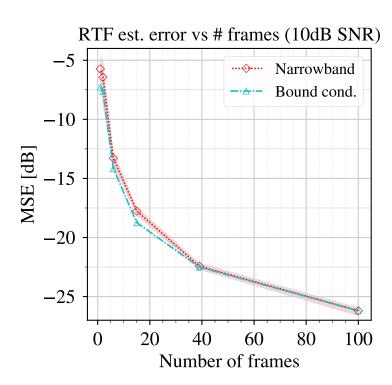
Experiments

Settings

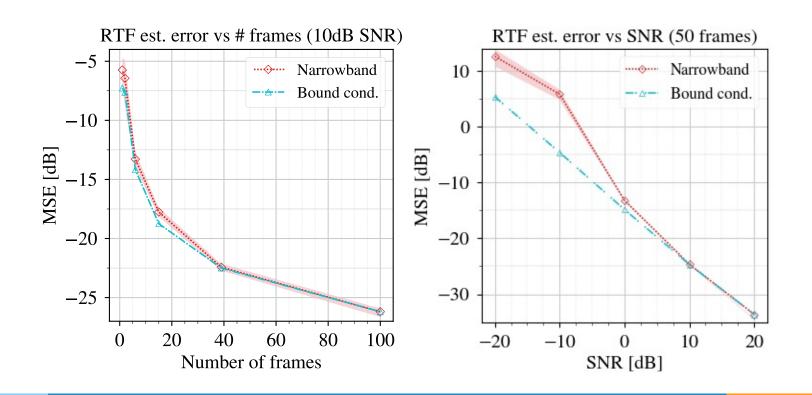
Performance metric

Mean-squared error (in dB) between actual and estimated RTFs

Experiments



Experiments



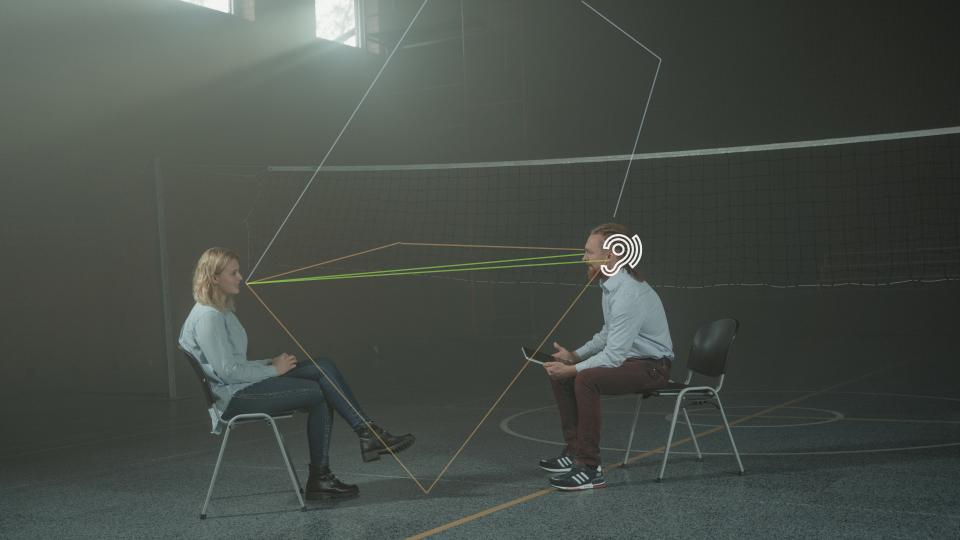
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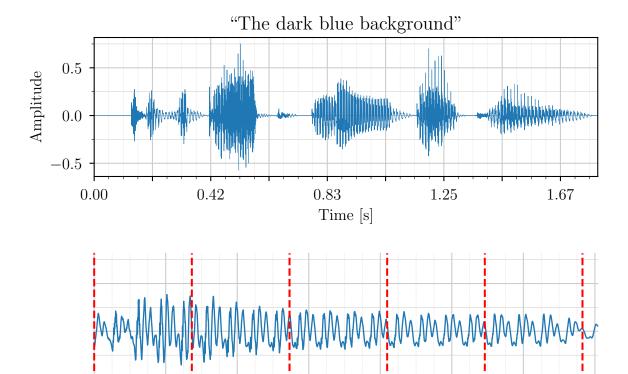
Cramér-Rao bound for acoustic transfer function estimation

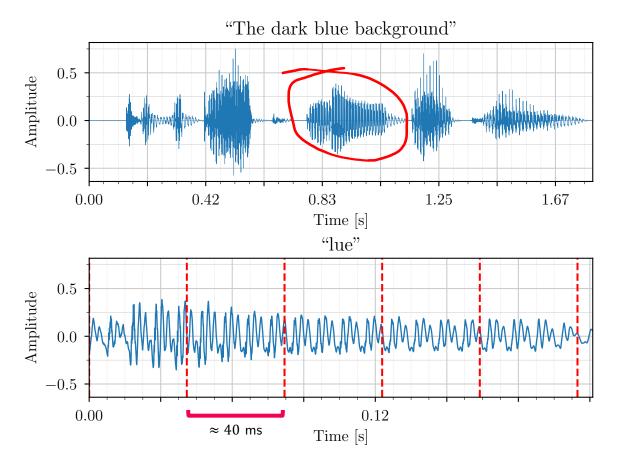
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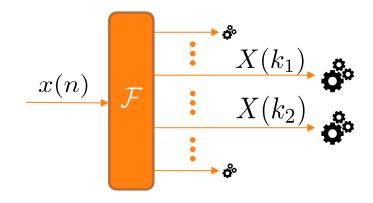
G. Bologni, R. Heusdens, and R. C. Hendriks "Harmonics to the rescue: Why voiced speech is not a WSS process", submitted for consideration at the 18th International Workshop on Acoustic Signal Enhancement (IWAENC 2024)

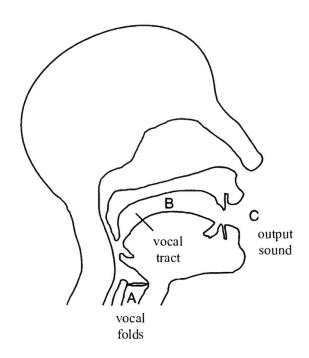
Inter-frequency correlations

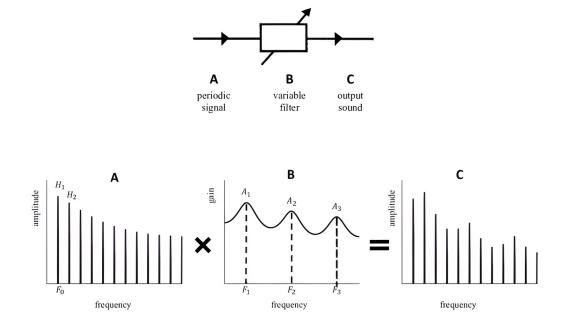
Signal processing algorithms (Wiener filter, MVDR) operate on **frequency-by-frequency** basis.

Frequency bins often assumed mutually uncorrelated.

Is assumption verified?



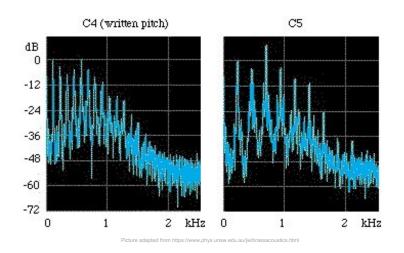




Inter-frequency correlations

Inter-frequency correlations found in

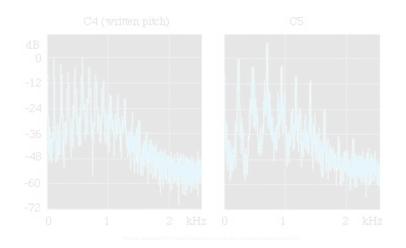
- 1. Speech, wind instruments
- 2. Windowed signals ("frequency leakage" effect)
- 3. Non-stationary signals



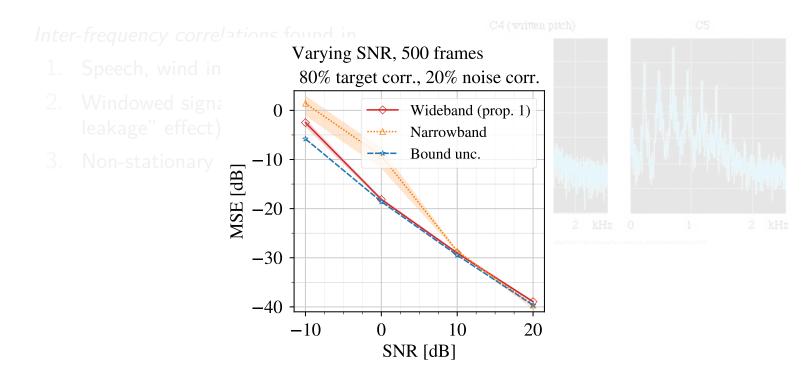
Can acoustic parameter estimation be improved by exploiting "hidden" correlations across frequencies?

Inter-frequency correlations found in

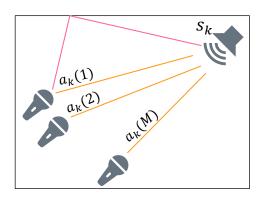
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Can acoustic parameter estimation be improved by exploiting "hidden" correlations across frequencies?



Channel estimation



Signal model – all frequencies

Noisy coefficients corresponding to single time frame l, at K different frequencies, can be stacked in a column as

$$\mathbf{x} = egin{bmatrix} \mathbf{x}_1 \ \mathbf{x}_2 \ dots \ \mathbf{x}_K \end{bmatrix} \in \mathbb{C}^{KM}.$$

Clean speech coefficients s and noise coefficients s can be obtained in the same way.

Signal model – all frequencies

The spatio-frequency correlation matrix R_x can then be expressed as

$$\mathbf{R}_{\mathbf{x}} = \mathbf{E} \begin{bmatrix} \mathbf{r}_{x}(1,1) & \mathbf{r}_{x}(1,2) & \cdots & \mathbf{r}_{x}(1,K) \\ \mathbf{r}_{x}(2,1) & \mathbf{r}_{x}(2,2) & \cdots & \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{r}_{x}(K,1) & \mathbf{r}_{x}(K,2) & \cdots & \mathbf{r}_{x}(K,K) \end{bmatrix} \in \mathbb{C}^{KM \times KM}, \quad (1)$$

where the bifrequency spatial correlation $\mathbf{r}_x(i,j)$ between noisy vectors at two arbitrary frequencies, i and j, is the spatial correlation matrix

$$[\mathbf{R}_{\mathbf{x}}]_{ij} = \mathbf{r}_x(i,j) = \mathbf{E}[\mathbf{x}_i \mathbf{x}_j] \in \mathbb{C}^{M \times M}.$$
 (2)

Signal model – second order

$$\mathbf{R}_{x} = \mathbf{E} \left[\mathbf{x} \mathbf{x}^{H} \right]$$

$$= \mathbf{E} \left[(\mathbf{A} \mathbf{s} + \mathbf{v}) (\mathbf{A} \mathbf{s} + \mathbf{v})^{H} \right]$$

$$= \mathbf{A} \mathbf{E} \left[\mathbf{s} \mathbf{s}^{H} \right] \mathbf{A}^{H} + \mathbf{E} \left[\mathbf{v} \mathbf{v}^{H} \right]$$

$$= \mathbf{A} \mathbf{R}_{s} \mathbf{A}^{H} + \mathbf{R}_{v}$$

$$= \mathbf{R}_{d} + \mathbf{R}_{v}$$

$$\mathbf{A} = \operatorname{diag}(\mathbf{a}) = \operatorname{diag}(a_{11}, \dots, a_{1M}, a_{21}, \dots a_{KM})$$

$$(5)$$

Signal model – second order

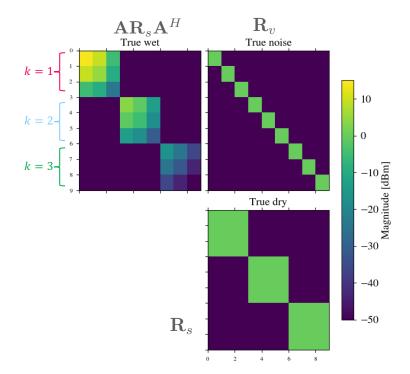
$$\mathbf{R}_{x} = \mathbf{E} [\mathbf{x} \mathbf{x}^{H}]$$
(1)

$$= \mathbf{E} [(\mathbf{A} \mathbf{s} + \mathbf{v})(\mathbf{A} \mathbf{s} + \mathbf{v})^{H}]$$
(2)

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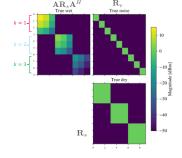
$$= \mathbf{A} \mathbf{R}_{s} \mathbf{A}^{H} + \mathbf{R}_{v}$$
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 (5)



RTF estimation – SVD direct

- We propose new RTF estimation algorithm that exploits inter-frequency correlations
- The algorithm is not 'optimal', but it sometimes achieve the CRB
 - O Q: what would it take for 'optimality'?



SVD-direct – key observation

When the number of frequencies is K=2, the clean covariance matrix can be decomposed as:

$$\mathbf{R_d} = \begin{bmatrix} \sigma_{s_1}^2 \mathbf{a_1} \mathbf{a_1}^H & \mathbf{E}[s_1 s_2^*] \mathbf{a_1} \mathbf{a_2}^H \\ \mathbf{E}[s_2 s_1^*] \mathbf{a_2} \mathbf{a_1}^H & \sigma_{s_2}^2 \mathbf{a_2} \mathbf{a_2}^H \end{bmatrix} = \begin{bmatrix} \mathbf{R_d^1} \\ \mathbf{R_d^2} \end{bmatrix}, \tag{1}$$

Notice that R_d^1 is rank-1 matrix.

Left principal singular vector of R_d^1 is a_1 . Applying SVD,

$$\mathbf{R_d^1} = \mathbf{SDV}^H \approx d \, \mathbf{a}_1 \mathbf{v}^H \tag{2}$$

SVD-direct – algorithm

1. Estimation of spatio-frequency covariance matrices: noisy $\hat{\mathbf{R}}_{\mathbf{x}}$, noise $\hat{\mathbf{R}}_{\mathbf{v}}$.

$$\hat{\mathbf{R}}_{\mathbf{x}} = \frac{1}{L} \sum_{l=1}^{L} \mathbf{x}(l) \mathbf{x}(l)^{H}$$
(1)

$$\hat{\mathbf{R}}_{\mathbf{v}} = \frac{1}{L} \sum_{l=1}^{L} \mathbf{v}(l) \mathbf{v}(l)^{H}$$
 (2)

2. Estimation of $\hat{\mathbf{R}}_{\mathbf{d}}$ from generalized eigenvector:

$$\hat{\mathbf{R}}_{\mathbf{x}}\mathbf{U} = \hat{\mathbf{R}}_{\mathbf{v}}\mathbf{U}\boldsymbol{\Lambda} \tag{3}$$

$$(\hat{\mathbf{R}}_{\mathbf{d}} + \hat{\mathbf{R}}_{\mathbf{v}})\mathbf{U} = \hat{\mathbf{R}}_{\mathbf{v}}\mathbf{U}\boldsymbol{\Lambda} \tag{4}$$

$$\hat{\mathbf{R}}_{\mathbf{d}} \approx \hat{\mathbf{R}}_{\mathbf{v}} \mathbf{U}_1 (\mathbf{\Lambda}_1 - \mathbf{I}) \mathbf{U}_1^{-1} \tag{5}$$

SVD-direct – algorithm

3. Partition in K "fat" $M \times KM$ blocks:

$$\hat{\mathbf{R}}_{\mathbf{d}} = \begin{bmatrix} \hat{\mathbf{R}}_{\mathbf{d}}^{(1)} \\ \hat{\mathbf{R}}_{\mathbf{d}}^{(2)} \\ \vdots \\ \hat{\mathbf{R}}_{\mathbf{d}}^{(K)} \end{bmatrix}. \tag{6}$$

4. SVD on individual subblocks

$$\hat{\mathbf{R}}_{\mathbf{d}}^{(k)} = \mathbf{P}^{(k)} \mathbf{D}^{(k)} \mathbf{Q}^{(k)}. \tag{7}$$

5. Rescale left principal singular vectors

$$\hat{\mathbf{a}}^{(k)} = \mathsf{Normalize}(\mathbf{p}_1^{(k)}).$$

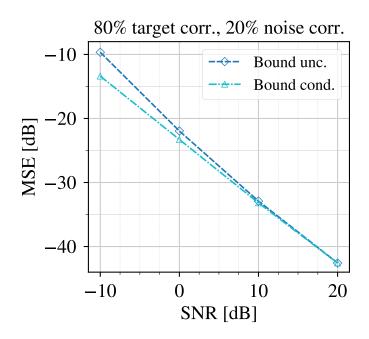
Two different CRBs

Conditional bound (see prev. slides)

True target received signal s is known

Unconditional bound

received signal s is unknown, but its firstand second-order statistics R_s are known



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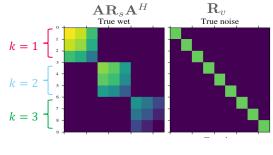
How to simulate data

- > To verify correctness of CRBs, we need access to the **true signal statistics**

Remember the signal model

$$\mathbf{R}_x = \mathbf{A}\mathbf{R}_s\mathbf{A}^H + \mathbf{R}_v$$

$$\mathbf{A} = \text{diag}(\mathbf{a}) = \text{diag}(a_{11}, \dots, a_{1M}, a_{21}, \dots a_{KM})$$
 (2)



Controlling correlations

Example for K=2 frequencies, M=2 sensors. Variable power.

$$\mathbf{R}_{v}^{\text{both}} = \begin{bmatrix} \sigma_{v_{11}}^{2} & \rho \sqrt{\sigma_{v_{11}}^{2} \sigma_{v_{12}}^{2}} & \cdots & \rho \sqrt{\sigma_{v_{11}}^{2} \sigma_{v_{22}}^{2}} \\ & \sigma_{v_{12}}^{2} & \cdots & \vdots \\ & & \sigma_{v_{21}}^{2} & \rho \sqrt{\sigma_{v_{21}}^{2} \sigma_{v_{22}}^{2}} \end{bmatrix}, \tag{1}$$

where $\sigma_{v_{ii}}^2 \sim \mathcal{U}(5e-4, 5e-1), \forall i = 1, ..., MK$ and $\rho \in [0, 1]$.

Experiments

M = 2 microphones

K = 5 frequency components,

Random RTF a/a_{ref} , real and imaginary parts drawn from Uniform(-1,1)

Covariance matrix R_x estimated over multiple snapshots

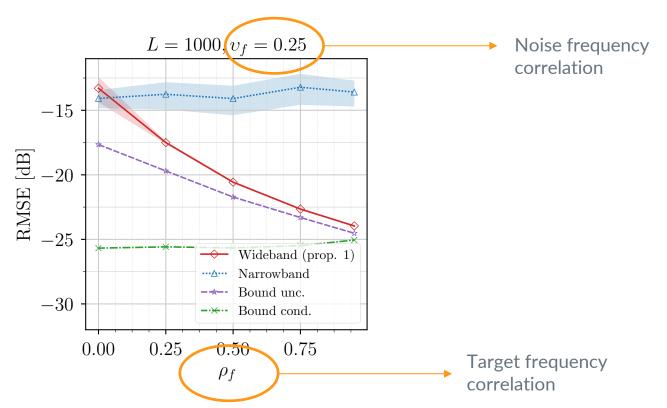
Oracle knowledge of R_v

Target s(l) and noise v(l) signals drawn from complex Gaussian distributions with given covariance matrices

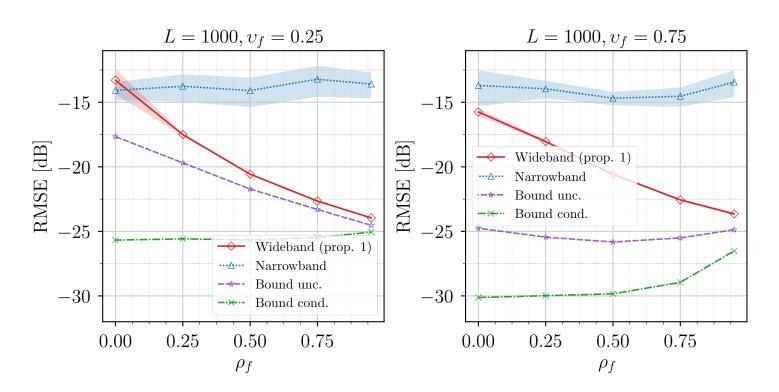
Performance metric

Mean-squared error (in dB) between actual and estimated RTFs

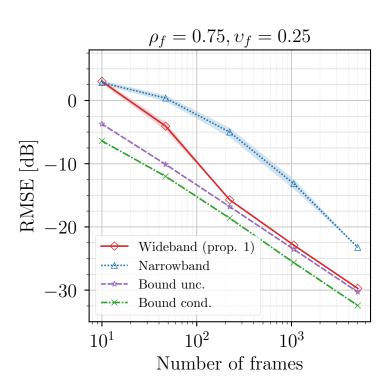
Synthetic data - varying target correlation



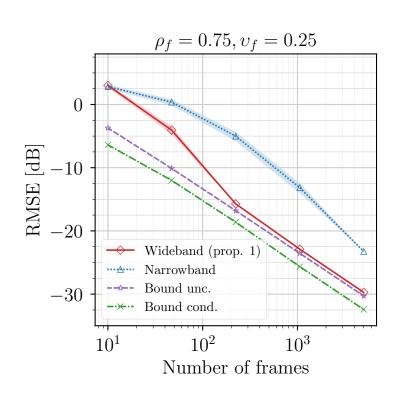
Synthetic data - varying target correlation

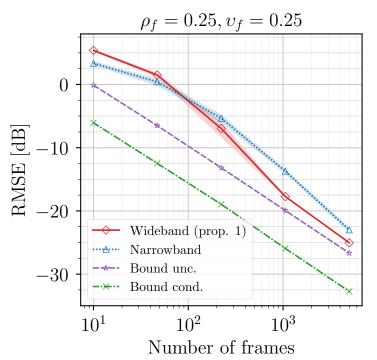


Synthetic data – varying number of frames

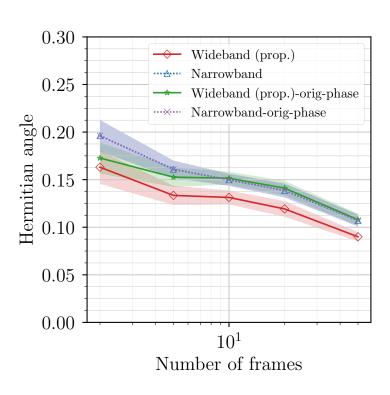


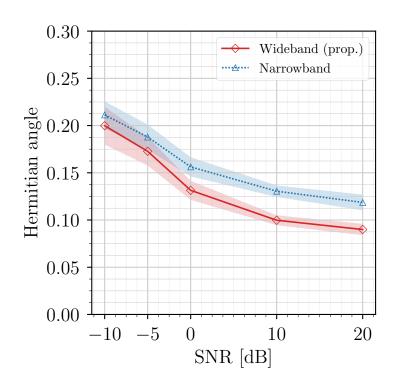
Synthetic data – varying number of frames

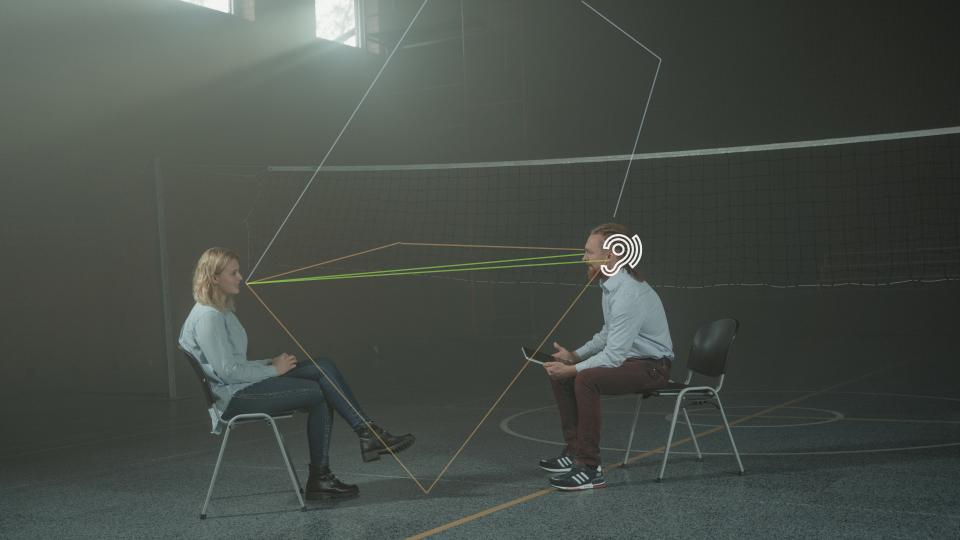




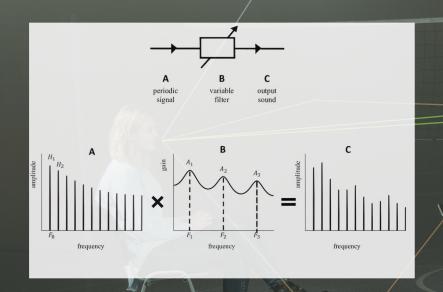
Real data

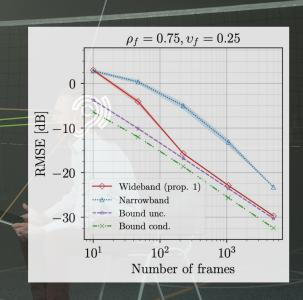




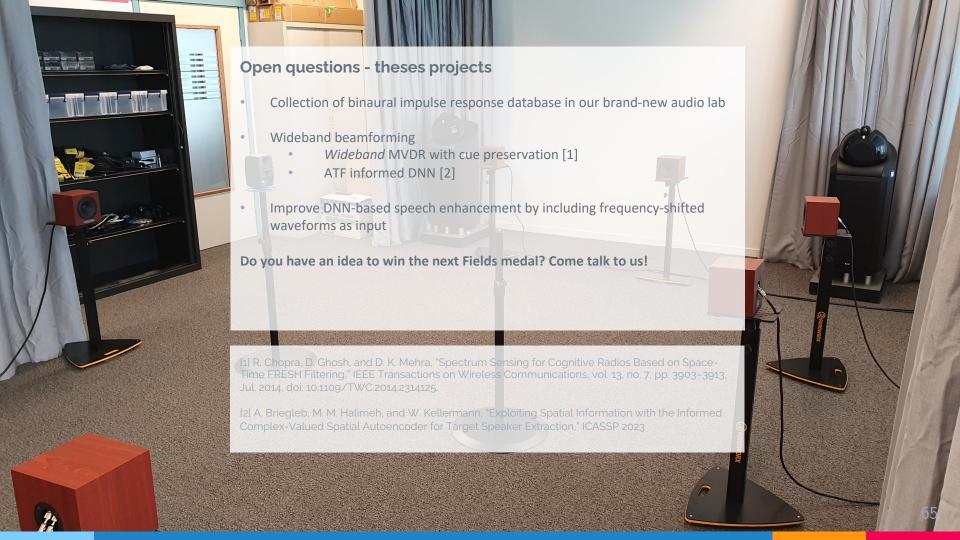












Recap

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